# Supplementary Materials: The Politics of Imperial Nostalgia

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### 1. Details of survey

Our data come from a three-round panel survey conducted by YouGov with a sample of adult residents of Britain (excluding Northern Ireland). The first round, fielded from October 3 to 19, 2022, included 4,069 respondents. YouGov employs quota sampling rather than probability sampling, but this sample was designed to be representative of the adult populations of England, Scotland, and Wales. Quotas based on age, gender, social class, region, and political attitudes were used, with post-stratification weighting applied to match population demographics. Our sample includes deliberate oversamples of rural residents by design, but all results are weighted to be representative of the adult populations of England, Wales, and Scotland at the time of survey administration.

The second round, conducted from October 13 to November 11, 2023, retained 2,522 participants from the first wave, and the third round, fielded from May 21 to June 10, 2024, included 2,169 of the original respondents, with 1,876 participating in all three waves. The first wave featured an oversample of rural residents, and weights were constructed for each survey wave using raking to align results with national population marginals.

To improve response quality, respondents who completed a survey wave in less than one-third of the median completion time were excluded: 117 respondents were excluded in wave 1 (median 24.7 minutes) and 12 in wave 2 (median 19.1 minutes), with none excluded in wave 3. Additionally, 36 respondents who were missing demographic data and six who had left the UK during the panel period were dropped. The final sample sizes for analysis were 3,884 for wave 1, 2,451 for wave 2, and 2,109 for wave 3.

## 2. Psychometrics of the scales

#### 2.1. Attitudes to empire

The imperial attitudes battery appears to measure a single coherent and reliable dimension of opinion. The scree plots from both survey rounds reveal a strong first eigenvalue, a second eigenvalue below one, and a pronounced "elbow" at the second eigenvalue, all of which support a unidimensional interpretation (see the tables and figures in this supplementary materials document). A one-dimensional confirmatory factor analysis fits reasonably well, with the CFI and SRMR metric better than the conventional metrics, but the RMSEA metric falling short (round 2: CFI = .996, RMSEA = .105, SRMR = .043; round 3: CFI = .997, RMSEA = .106, SRMR = .040). In addition, all items have strong positive standardised loadings in the .74-.89 range. The battery is internally consistent across both rounds ( $\alpha$  = .91 and .92), showing inter-item reliability. The repeated measurements allow us to assess the test-retest reliability, which is also strong (r = .87).

#### 2.2. Emotions to empire

The imperial emotions battery displays a more complex structure than imperial attitudes, with evidence for distinct positive and negative dimensions. Such a finding is consistent with classic research on the latent structure of emotional response (?). Specifically, scree plots reveal that the first two eigenvalues are elevated (both greater than 2), with a distinct "elbow" at the third. Exploratory factor analyses reveal a clear separation, with the items tapping negative and positive emotions loading on separate factors. Confirmatory factor analyses show that the two-dimensional, negative vs. positive imperial emotion model fits well (CFI = .999, RMSEA = .053, SRMR = .0388; r3: CFI = .999, RMSEA = .054, SRMR = .039) while a one-dimensional model does not (r2: CFI = .972, RMSEA = .296, SRMR = .240; r3: CFI = .980, RMSEA = .261, SRMR = .202). The two dimensions exhibit a moderate negative correlation (r = -.51 & -.60).

While these results support the use of separate negative and positive emotions to empire scales, creating a single differenced measure (i.e., positive emotions scale – negative emotions scale) produces a scale with greater test-retest reliability (r = .86 vs. .81) and a stronger correlation with imperial attitudes ( $r = .87 \cdot .88 \text{ vs. } .74 \cdot .78$ ). Given the strong correlation that imperial attitudes exhibits with the net imperial emotions scale, as well as the stronger correlation it shows with other measures of cultural conservatism (compared with net imperial emotions; results in online supplementary materials), we use the imperial attitudes scale as our measure of imperial nostalgia in the remainder of the paper.

**Table S1.** Summary of imperial attitude positions

	disagree	unsure	agree	net % nostalgic
harm vs. good	29.2	34.7	36.1	-6.9
civilising effect	19.1	34.1	46.9	27.8
advanced humanity	20.3	40	39.8	19.5
responsible atrocities	10.8	26.3	62.9	-52.1
wish empire	54.0	29.6	16.4	-37.7
golden age	29.2	36.4	34.3	5.1
shameful period	32.3	32.5	35.2	-3.0

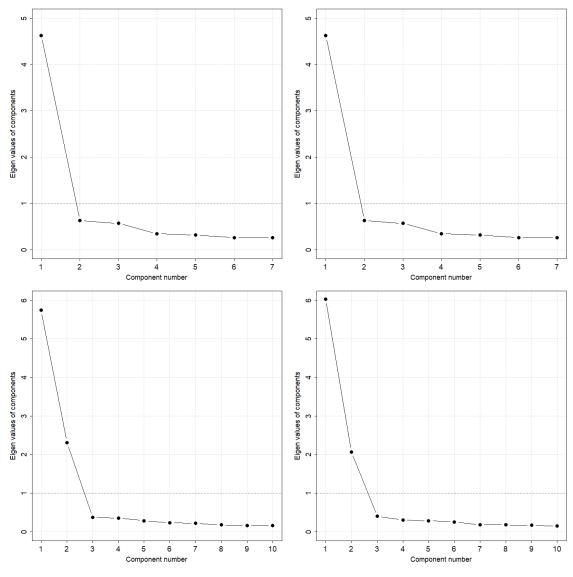


Figure S1. Scree plots, imperial nostalgia

Notes: Each panel shows a scree plot displaying the eigenvalues of the correlation matrix for the items. Top row: imperial attitudes round 2 (left), round 3 (right); bottom row: imperial emotions, round 2 (left), round 3 (right)

**Table S2.** CFA, imperial attitudes, round 3 survey

Factor Loadings		
	Estimate	Std. Err.
imp att q1	0.88	0.01
imp att q2	0.89	0.01
imp att q3	0.86	0.01
imp att q4	0.77	0.01
imp att q5	0.74	0.01
imp att q6	0.85	0.01
imp att q7	0.88	0.01
Latent Variances		
Imperial attitudes	1.00+	
Fit Indices		
N	2109	
$\chi^2$	339.92	
CFI	1.00	
RMSEA	0.11	
SRMR	0.04	
Scaled $\chi^2(df)$	1002.95(14)	

<sup>+</sup>Fixed parameter

Table S3. Correlations between imperial nostalgia scales

	IA r2	IA r3	IE+ r2	IE+ r3	IE- r2	IE- r3	IED r2
IA r3	.87						
IE+r2	.78	.74					
IE+r3	.77	.81	.81				
IE- r2	76	73	55	59			
IE-r3	72	78	57	64	.81		
IED r2	.87	.83	.88	.79	88	78	
IED r3	.82	.88	.76	.90	77	91	.86

IA: Imperial attitudes; IE+ Imperial emotions, positive; IE- Imperial emotions, negative; IED Imperial emotions, difference; r2,r3: round 2,3

**Table S4.** Exploratory factor analysis

	1	2	3	4	5	6	7	8
imperial nost. 1	0.92							
imperial nost. 2	0.84							
imperial nost. 3	0.84							
imperial nost. 4	0.71							
imperial nost. 5	0.66							
imperial nost. 6	0.76							
imperial nost. 7	0.92							
immigr. eco.						0.80		
immigr. cult.						0.76		
eco. values 1							0.80	
eco. values 2		0.39					0.38	
eco. values 3		0.30					0.70	
eco. values 4		0.46					0.55	
auth. values 1					0.65			
auth. values 2					0.42			
auth. values 3					0.62			
auth. values 4					0.65			
sexism 1								0.82
sexism 2								0.85
sexism 3								0.79
sup. democ. 1			0.73					
sup. democ. 2			0.73					
sup. democ. 3			0.56					
sup. democ. 4			0.80					
sup. democ. 5			0.73					
sup. democ. 6			0.46					
sup. democ. 7			0.70					
sup. democ. 8			0.70					
satis. democ.		-0.41			0.48			
trust MPs		-0.69						
gov. approval		-0.45			0.42			
left-right					~		-0.42	
EU indepen.						-0.40		
populism 1		0.92						
populism 2		0.84						
populism 3		0.92						
populism 4		0.57						
chauv. nation. 1				0.77				
chauv. nation. 2				0.76				
chauv. nation. 3				0.92				
nation. pride 1				0.67				
nation. pride 2				0.69				
nation. pride 3				0.64				
F								

Exploratory factor analysis (EFA) with eight dimensions (as indicated with parallel analysis), using minimum residual estimation and promax rotation. Based on pairwise polyserial correlations. Loadings < |0.30| suppressed for clarity. Most items are from round 2, except national pride and chauvinistic nationalism (round 3).

**Table S5.** EFA inter-factor correlations

	F1	F2	F3	F4	F5	F6	F7	F8
F1	1.00	-0.09	-0.41	0.60	0.63	-0.42	-0.55	0.51
F2	-0.09	1.00	-0.08	-0.31	-0.04	-0.41	0.41	0.13
F3	-0.41	-0.08	1.00	-0.19	-0.44	0.47	0.27	-0.43
F4	0.60	-0.31	-0.19	1.00	0.58	-0.05	-0.47	0.25
F5	0.63	-0.04	-0.44	0.58	1.00	-0.31	-0.46	0.44
F6	-0.42	-0.41	0.47	-0.05	-0.31	1.00	0.10	-0.47
F7	-0.55	0.41	0.27	-0.47	-0.46	0.10	1.00	-0.25
F8	0.51	0.13	-0.43	0.25	0.44	-0.47	-0.25	1.00

Inter-factor correlations from the Exploratory factor analysis (EFA). The imperial nostagia factor is F1.

Figure S2. Correlations between measures of imperial nostalgia and other attitudinal variables

	Imperial Attitudes	Imperial Emotions
Immigration support	-0.58	-0.55
Authoritarian values -	0.57	0.52
National chauvinism -	0.56	0.55
EU independence	0.56	0.54
National pride	0.53	0.51
Hostile sexism -	0.50	0.47
Authoritarian child rearing values	0.42	0.41
Left-economic values	-0.38	-0.36
Democratic support	-0.33	-0.29
British identity	0.33	0.32
English identity -	0.33	0.30
Satisfaction with _ democracy	0.26	0.27
Life better 50 years _ ago	0.16	0.11
Trust in politicians	0.15	0.19
External efficacy -	0.10	0.11
Populist attitudes -	-0.01	-0.04

The heatmap shows the bivariate correlations between the two measures of imperial nostalgia and other attitudinal variables we collected. Darker blue cells indicate correlations with larger absolute values. Wave 2 correlations are reported, except for national chauvinism, national pride and authoritarian child-rearing (all in wave 3) and nostalgic deprivation and life better 50 years ago (measured in wave 1, correlated with nostalgia in wave 2

**Table S6.** OLS regressions of party evaluations, round 2

	Cons	Lab	LDem	Reform	Green	SNP
Imperial nostalgia	.42*	22*	12	.49*	54*	75*
	(.09)	(.09)	(.09)	(.08)	(.09)	(.36)
Left-eco.values	-1.04*	.97*	.15	.01	.39*	.53
	(.09)	(.10)	(.09)	(.08)	(.09)	(.34)
Authoritarian values	.85*	.19	.22*	16	29*	59
	(.10)	(.11)	(.10)	(.10)	(.10)	(.40)
Immigration support	.29	.87*	.97*	$-1.05^*$	.49*	.20
	(.16)	(.18)	(.16)	(.16)	(.16)	(.66)
Hostile sexism	16	$50^{*}$	$30^{*}$	.17*	$40^{*}$	.63
	(.08)	(.09)	(.09)	(.08)	(80.)	(.36)
Populist attitudes	66*	13	20*	.44*	.03	09
	(.09)	(.09)	(.09)	(.08)	(.09)	(.34)
English identity	.05*	.04	.02	.03	.00	16
	(.02)	(.02)	(.02)	(.02)	(.02)	(.09)
EU independence	.26*	34*	23*	.23*	$19^*$	$30^{*}$
	(.02)	(.02)	(.02)	(.02)	(.02)	(.11)
$R^2$	.47	.43	.26	.35	.38	.41
N	2029	2028	2011	1897	2011	242

<sup>\*</sup>p < 0.05. OLS regression models.

**Table S7.** OLS regressions of pro- (vs. anti-) empire candidates

	Model 1	Model 2
Intercept	.53*	.53*
•	(.04)	(.04)
Imperial nostalgia	.17*	.16*
	(.02)	(.03)
Authoritarian values	00	.02
	(.03)	(.03)
Conservative party evaluations	00	
	(.01)	
Labour party evaluations	.01	
	(.01)	
Liberal Democrat evaluations	00	
	(.01)	
Reform party evaluations	.01	
	(.01)	
Green party evaluations	01	
	(.01)	
Vote intention for left (vs. right) party	I	.03
		(.05)
$R^2$	.16	.11
N	1268	1017

 $<sup>^*</sup>p < 0.05$ . OLS regression models, with standard errors clustered by respondent in parentheses. The outcome is 1 when respondents were presented with a pro- vs anti-empire pair of candidates and selected the pro-empire one; and 0 if they selected the anti-empire candidate.

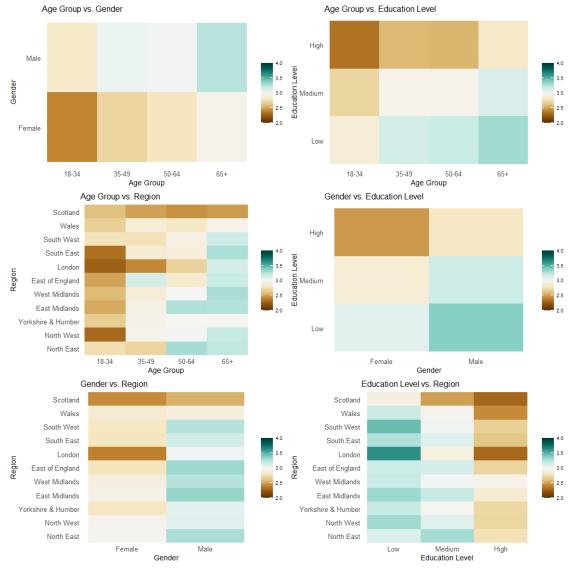


Figure S3. Imperial nostalgia by 2-way demographics

Results of linear models regressing the composite mean imperial attitudes scale on each pair of the following demographic variables: age groups, gender, region, and education level, with interaction terms included. The cells show predicted effects of these linear models, which are obtained using data pooled across survey rounds and are weighted.

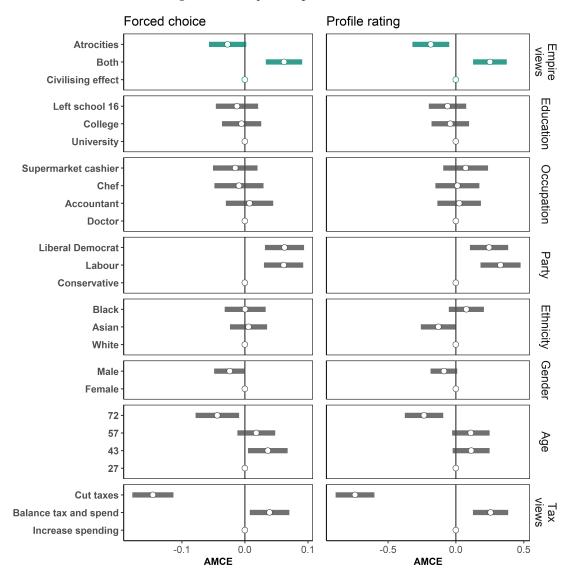


Figure S4. Conjoint experiment, AMCEs

Results of conjoint experiment - average marginal component effects (AMCEs)

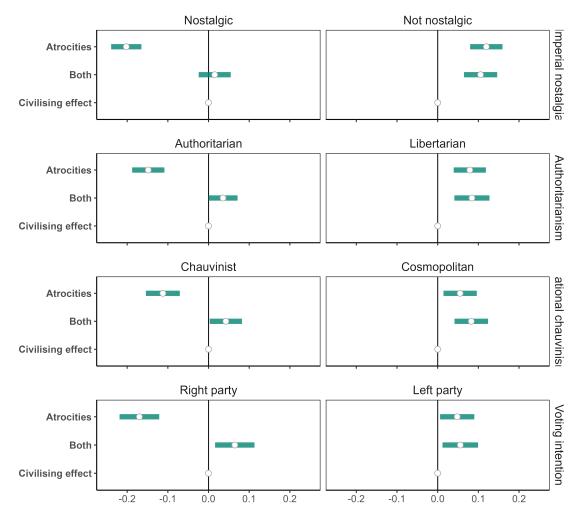


Figure S5. Conjoint experiment, AMCEs by attitudinal subgroups

Results of conjoint experiment (average marginal component effects) when splitting the sample by the median of imperial nostalgia, authoritarianism, and national chauvinism, and by right vs. left party preference. Results from the forced MP choice question used.

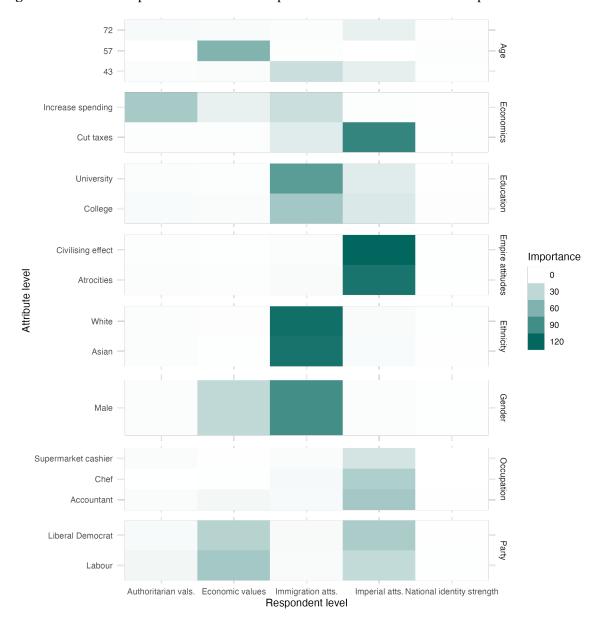


Figure S6. Variable importance matrix from separate random forest models on separate attribute levels.

Notes: This figure presents variable importance scores for respondent-level covariates (columns) in predicting the distribution of (individual-level) effects for each attribute level (rows). Higher values (darker shading) indicate variables' greater importance in predicting the estimated effect distribution. The individual level effects are derived from Bayesian Additive Regression Trees (BARTs), and then the variable importance is derived from random forests.